

## Generalizations of Tukey distributions

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### Abstract

The Generalized Lambda Distribution (GLD) is a four-parameter generalization of Tukey's Lambda family. Several methods for estimating the parameters of the GLD have been reported in the literature, but the most popular one is the moment-method matching proposed by Ramberg and Schmeiser (1974). One of the drawbacks of the GLD is that the shape parameter also determines skewness. It seems reasonable that there should be three linear parameters determining position, scale, and skewness and two parameters determining the shapes of the two tails. This suggests a natural generalization of the GLD to give a five-parameter lambda distribution (FPLD).

**Keywords:** Tukey distribution, moment method, lambda distribution.

### 1. Foreword

Varied empirical distributions have been reported in statistical literature. The Pearson system (Pearson (1948), the Johnson system (Hahn and Shapiro (1967)), the Burr distribution (Burr (1973)) and Tukey distribution (1960) belong to the most important empirical distributions.

In order to derive probability distribution function (PDF) one can follow two approaches.

Pearson (1895) specified differential equation as the basis to derive the family of PDF

$$\frac{df(x)}{dx} = \frac{(x-c)f(x)}{a_0 + a_1x + a_2x^2} \quad (1)$$

where  $f(x)$  is a density function. The constant coefficients entering the distribution can be described in terms of the first four moments ( when they exist).

Solutions of eq. (1) are classified according to the character of roots of equation  $a_0 + a_1x + a_2x^2 = 0$ , where

$$a_0 + a_1x + a_2x^2 = a_2 \left( x - \frac{(a_1 - \sqrt{a_1^2 - 4a_0a_2})}{2a_2} \right) \left( x - \frac{(a_1 + \sqrt{a_1^2 - 4a_0a_2})}{2a_2} \right). \quad (2)$$

Edgeworth (1898) presented another approach, the so called method of translation which is based on two functions. D'Addaro (1949) called these functions the probability generating function (PGF) and transformation function (TF).

The Johnson system (1949) provides wide variety of distribution shapes as wide as that of Pearson system. In Edgeworth system, the normal distribution plays a role of PGF and the logarithm of income – the role of TF.

### 2. Generalizations of Tukey-lambda distributions

In most cases a continuous probability distribution is defined by means the cumulative distribution function (CDF) or probability density function (PDF); alternatively, it can be defined in terms of percentile functions. This is the case for

Tukey-lambda distribution we look for the distribution parameters on the basis of the percentile function which is the inverse of the cumulative distribution function.

A particular example of the percentile function is Tukey's  $\lambda$  function

$$R(p) = \frac{p^\lambda - (1-p)^\lambda}{\lambda}, \quad 0 \leq p \leq 1, \quad \lambda \neq 0. \tag{3}$$

Ramberg and Schmeiser (1974) generalized Tukey's  $\lambda$  distribution to a four-parameter distribution defined by the percentile function

$$R(p) = \lambda_1 + \frac{p^{\lambda_3} - (1-p)^{\lambda_4}}{\lambda_2}, \quad 0 \leq p \leq 1, \tag{4}$$

where  $\lambda_1$  is a location parameter,  $\lambda_2$  is a scale parameter,  $\lambda_3$  is a skewness parameter and  $\lambda_4$  is a kurtosis one.

The density function is defined by

$$f(x) = \frac{1}{\frac{dR(p)}{dp}} = \frac{\lambda_2}{\lambda_3 p^{\lambda_3-1} + \lambda_4 (1-p)^{\lambda_4-1}}, \quad 0 \leq p \leq 1. \tag{5}$$

Some suggestions concerning G $\lambda$ D appeared in literature, namely that G $\lambda$ D is not suitable for description of the asymmetrical distributions. This is due to the fact that the shape parameter determines also skewness, so there should be rather three linear parameters determining position, scale, and skewness together with two parameters determining the shapes of the two tails. This suggests a natural generalization of the G $\lambda$ D to give a five-parameter lambda distribution (FPLD).

The five-parameter version of generalized lambda distribution was proposed by Gilchrist (2000). The quantile function is given by

$$Q(\lambda) = \lambda_1 + \frac{\lambda_2}{2} \left( (1-\lambda_3) \frac{p^{\lambda_4} - 1}{\lambda_4} - (1+\lambda_3) \frac{(1-p)^{\lambda_5} - 1}{\lambda_5} \right) \tag{6}$$

where  $0 \leq p \leq 1$ ,  $\lambda_2 \geq 0$ ,  $\lambda_3 \in \langle -1, 1 \rangle$ .

### 3. Comparison of G $\lambda$ D and FPLD approximations for asymmetric distribution

Several methods for estimating the parameters of the G $\lambda$ D exist.

- 1) The oldest method, the so called moment-matching method was proposed by Ramberg and Schmeiser (1979). In this method we equate the mean, the variance, and the third and fourth moments of G $\lambda$ D to the corresponding mean  $\mu^*$ , variance  $(\sigma^*)^2$ , skewness  $\beta_3^*$ , and kurtosis  $\beta_4^*$  of the sample and then compute the parameters  $\lambda$  from these four equations.
- 2) Öztürk and Dale (1985) proposed the least squares method. In this method one finds the values of  $\lambda$  for which the differences between the observed and predicted order statistics are as small as possible.
- 3) The starship method was proposed by King and MacGillivray (1999). First, we select a region in four-dimensional space that covers the range of the four parameters  $\lambda$  and cover it by a four-dimensional rectangular grid. Next we evaluate the grid points by performing a goodness-of-fit test on the corresponding distributions. If the test is satisfied, we stop procedure; otherwise we continue with the next point in the grid.

Lakhany and Maussner (2000) compared the results of the above mentioned estimation methods. It follows from their analysis that the method proposed by

Öztürk and Dale is one of the best. For that reason we use only this method. In more detail, the procedure based on the least squares method can be described as follows. Let  $x_i, i = 1, \dots, n$  denote the  $i$ th order statistic of data which is to be represented by the quantile function  $R(u)$  and let  $U_i, i = 1, \dots, n$  denote the order statistic of the corresponding uniformly distributed random variable. The method consists in determining the values of  $\lambda$  for which the differences between the observed and predicted order statistics are as small as possible. So, in the case of G $\lambda$ D we must minimize the function

$$G(\lambda) = \sum_{i=1}^n \left( x_{(i)} - \lambda_1 - \frac{Z_i}{\lambda_2} \right)^2 \tag{7}$$

where

$$\begin{aligned} Z_i &= \frac{1}{\lambda_3} (EU_{(i)}^{\lambda_3} - 1) - \frac{1}{\lambda_4} (E(1 - U_{(i)})^{\lambda_4} - 1) \\ EU_{(i)}^{\lambda_3} &= \frac{\Gamma(n+1)\Gamma(i + \lambda_3)}{\Gamma(i)\Gamma(n + \lambda_3 + 1)} \\ E(1 - U_{(i)})^{\lambda_4} &= \frac{\Gamma(n+1)\Gamma(n - i + \lambda_4 + 1)}{\Gamma(n - i + 1)\Gamma(n + \lambda_4 + 1)} \end{aligned} \tag{8}$$

On the other hand, in the case of FPLD we minimize the function

$$G(\lambda) = \sum_{i=1}^n (x_{(i)} - Z_{(i)})^2 \tag{9}$$

where

$$Z_{(i)} = \lambda_1 + \frac{(1 - \lambda_3)\lambda_2}{2\lambda_4} \left( \frac{\Gamma(n+1)\Gamma(i + \lambda_4)}{\Gamma(i)\Gamma(n + 1 + \lambda_4)} - 1 \right) + \frac{(1 + \lambda_3)\lambda_2}{2\lambda_5} \left( 1 - \frac{\Gamma(n+1)\Gamma(n + 1 - i + \lambda_5)}{\Gamma(n + 1 - i)\Gamma(n + 1 + \lambda_5)} \right) \tag{10}$$

In both cases we minimize functions with the help of means of Nelder-Mead method.

We have drawn random samples of size  $n = 400$  for the chosen asymmetric continuous distributions. We calculated the parameters  $\lambda_i$  of G $\lambda$ D and FPLD (Table 1) for every distribution. As a measure of quantitative closeness of two statistical distribution we used, following Taristano (2010), the Tchebycheff metric, which determine the maximum absolute value of the difference between the theoretical quantiles  $R(p)$  and fitted quantiles  $R(p, \lambda)$ :

$$\max_{1 \leq i \leq 39} |R(p_i) - R(p_i, \lambda)| \quad \text{where } p_i = \frac{i}{40}, i = 1, 2, \dots, 39. \tag{11}$$

Comparing distances for four- and five-parameters functions, calculated on the basis of eq. (11), we always obtained smaller distance for FPLD. This means that the FPLD provides a better fit than G $\lambda$ D.

Table 1. Comparison of fit  $G\lambda D$  i FPLD for asymmetric probabilistic distribution.

Model		$\lambda_1$	$\lambda_2$	$\lambda_3$	$\lambda_4$	$\lambda_5$	$\max_{1 \leq i \leq 39}  R(p_i) - R(p_i, \lambda) $
Beta[3,6]	$G\lambda D$	0.25263	2.05295	0.10853	0.38998		0.0154
	FPLD	0.29944	0.23902	0.23451	0.29279	0.30638	0.0120
Chi-square[4]	$G\lambda D$	1.16817	0.05299	0.01038	0.18126		0.0553
	FPLD	2.37635	3.99807	0.27658	0.71823	0.09292	0.0389
Gamma[2,2]	$G\lambda D$	1.79032	0.07079	0.02577	0.23227		0.7068
	FPLD	2.9177	3.81181	0.26153	0.4982	0.13424	0.6090
Lognormal[0,1]	$G\lambda D$	0.429	-0.21452	-0.02024	-0.1977		0.1349
	FPLD	2.62411	4.74602	-0.91609	1.89853	-0.5331	0.0517

Source: own calculations.

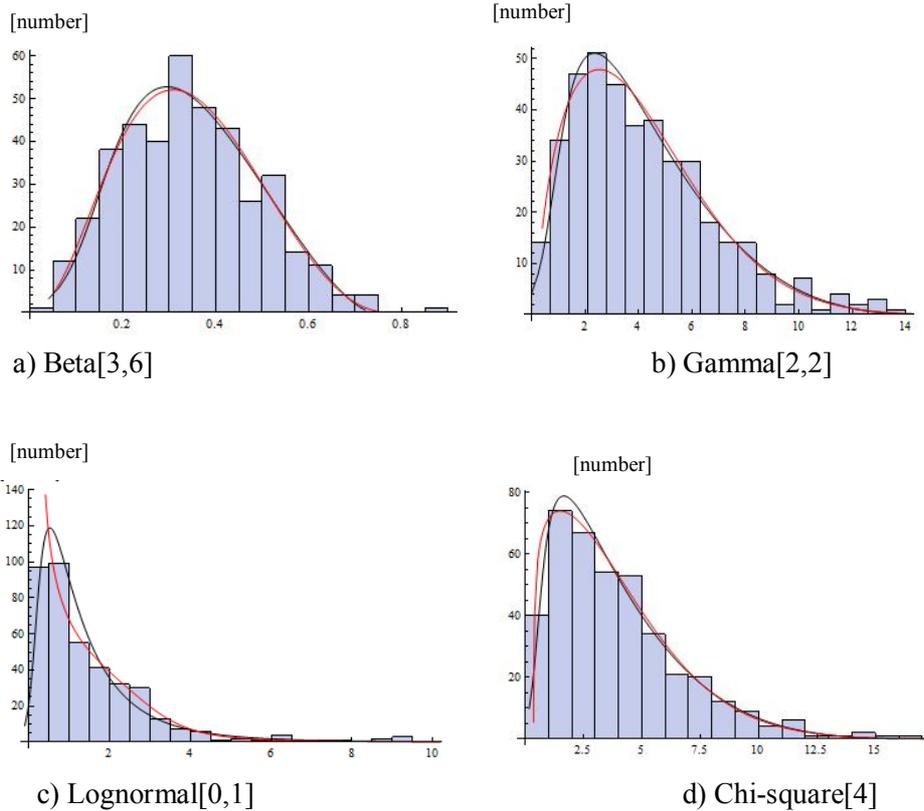


Figure. 1. Histograms of random samples and estimated functions  $G\lambda D$  (black curve) and FPLD (red curve).

#### 4. Conclusion

The generalized Tukey distribution permits to obtain a wide variety of curves, the simplest examples being presented on Fig. 1. Due to the large flexibility of this distribution, the  $G\lambda D$  finds many applications in the cases when the real distribution is not known.

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